

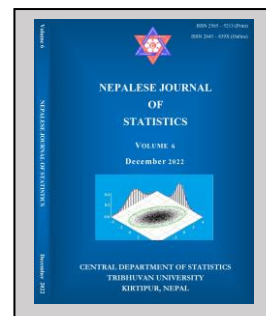
Supplementary Material

Inverse Exponentiated Odd Lomax Exponential
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SUPPLEMENTARY MATERIAL

(1) Formation of the Inverse Exponentiated Odd Lomax Exponential Distribution (IEOLE)

The cumulative distribution function of inverse exponential distribution is given as

$$G(x) = e^{-\alpha/x}, \quad \bar{G}(x) = 1 - G(x) = 1 - e^{-\alpha/x}, \quad x > 0, \alpha > 0$$

$$W(x) = \left[\frac{G(x)}{\bar{G}(x)} \right]^\theta = \left[\frac{e^{-\alpha/x}}{1 - e^{-\alpha/x}} \right]^\theta = (e^{\alpha/x} - 1)^{-\theta}$$

$$F(x) = \int_0^{W(x)} r(t) dt$$

Here, $r(t)$ is the generator which uses the probability density function (PDF) of the Lomax distribution (Chakraborty, 2019) and hence, the PDF of the Lomax distribution as a generator is provided by

$$r(t) = \frac{\lambda}{\delta} \left\{ 1 + \left(\frac{t}{\delta} \right) \right\}^{-(\lambda+1)}; \quad t > 0, \lambda > 0, \delta > 0$$

(A) Derivation of CDF:

$$F(x, \alpha, \lambda, \theta, \delta) = \int_0^{(e^{\alpha/x} - 1)^{-\theta}} \frac{\lambda}{\delta} \left\{ 1 + \frac{t}{\delta} \right\}^{-(\lambda+1)} dt; \quad = \frac{\lambda}{\delta} \int_0^{(e^{\alpha/x} - 1)^{-\theta}} \left\{ 1 + \frac{t}{\delta} \right\}^{-(\lambda+1)} dt;$$

$$\text{Let } v = 1 + \frac{t}{\delta} \Rightarrow \frac{dv}{dt} = \frac{1}{\delta} \Rightarrow dt = \delta dv$$

$$\text{Now, } \int v^{-(\lambda+1)} \delta dv = \frac{\delta \cdot v^{-(\lambda+1)+1}}{-(\lambda+1)+1} = \frac{\delta \cdot v^{-\lambda}}{-\lambda} = \frac{\delta \cdot (1+t/\delta)^{-\lambda}}{-\lambda}$$

$$F(x, \alpha, \lambda, \theta, \delta) = \frac{\lambda}{\delta} \int_0^{(e^{\alpha/x}-1)^{-\theta}} \left\{1 + \frac{t}{\delta}\right\}^{-(\lambda+1)} dt; = \frac{\lambda}{\delta} \left[\frac{\delta(1+t/\delta)^{-\lambda}}{-\lambda} \right]$$

$$= - \left[\left\{1 + \frac{(e^{\alpha/x}-1)^{-\theta}}{\delta}\right\}^{-\lambda} - 1 \right] = 1 - \left\{1 + \frac{(e^{\alpha/x}-1)^{-\theta}}{\delta}\right\}^{-\lambda} = 1 - \left\{1 + \frac{1}{\delta} (e^{\alpha/x}-1)^{-\theta}\right\}^{-\lambda}$$

(2) Derivation of Probability density function of the Inverse Exponentiated Odd Lomax Exponential Distribution (IEOLE)

$$f(x) = \frac{dF(x)}{dx} = \frac{d \left[1 - \left\{1 + \frac{1}{\delta} (e^{\alpha/x}-1)^{-\theta}\right\}^{-\lambda} \right]}{dx} = - \frac{d \left\{1 + \frac{1}{\delta} (e^{\alpha/x}-1)^{-\theta}\right\}^{-\lambda}}{dx}$$

$$= \frac{\lambda}{\delta} \left\{1 + \frac{1}{\delta} (e^{\alpha/x}-1)^{-\theta}\right\}^{-(\lambda+1)} \times (-\theta) (e^{\alpha/x}-1)^{-\theta-1} \frac{de^{\alpha/x}}{dx}$$

$$= \lambda \left\{1 + \frac{1}{\delta} (e^{\alpha/x}-1)^{-\theta}\right\}^{-\lambda-1} \frac{1}{\delta} \frac{d(e^{\alpha/x}-1)^{-\theta}}{dx}$$

$$= - \frac{\lambda\theta}{\delta} \left\{1 + \frac{1}{\delta} (e^{\alpha/x}-1)^{-\theta}\right\}^{-(\lambda+1)} (e^{\alpha/x}-1)^{-(\theta+1)} \alpha e^{\alpha/x} (-1/x^2)$$

$$= \frac{\lambda\theta\alpha}{\delta} x^{-2} e^{\alpha/x} (e^{\alpha/x}-1)^{-(\theta+1)} \left[1 + \frac{1}{\delta} (e^{\alpha/x}-1)^{-\theta}\right]^{-(\lambda+1)}$$

(3) Verification of Probability density function of the Inverse Exponentiated Odd Lomax Exponential Distribution (IEOLE)

To verify $\int_0^\infty f(x) dx = 1$

$$\text{Now, definite integral} = \int_0^\infty \frac{\lambda\theta\alpha}{\delta} x^{-2} e^{\alpha/x} (e^{\alpha/x}-1)^{-(\theta+1)} \left[1 + \frac{1}{\delta} (e^{\alpha/x}-1)^{-\theta}\right]^{-(\lambda+1)} dx$$

$$\text{Let } 1 + \frac{1}{\delta} (e^{\alpha/x}-1)^{-\theta} = v \Rightarrow \frac{dv}{dx} = \frac{\alpha\theta}{\delta} \frac{e^{\alpha/x}}{x^2} (e^{\alpha/x}-1)^{-(\theta+1)} \Rightarrow dv = \frac{\alpha\theta}{\delta} \frac{e^{\alpha/x}}{x^2} (e^{\alpha/x}-1)^{-(\theta+1)} dx$$

$$\text{Now, indefinite integral} = \lambda \int [v]^{-(\lambda+1)} dv = \frac{\lambda(v)^{-(\lambda+1)+1}}{-(\lambda+1)+1} = \frac{\lambda(v)^{-\lambda}}{-\lambda} = -(v)^{-\lambda} =$$

$$-\left(1 + \frac{1}{\delta} (e^{\alpha/x}-1)^{-\theta}\right)^{-\lambda}$$

$$\therefore \text{Definite integral} = \left[-\left(1 + \frac{1}{\delta} (e^{\alpha/x}-1)^{-\theta}\right)^{-\lambda} \right]_0^\infty$$

$$\begin{aligned}
 &= \left[-\left(1 + \frac{1}{\delta} \{e^{\alpha/x} - 1\}^{-\theta}\right)^{-\lambda} \right]_0^{\infty} = -\left(1 + \frac{1}{\delta} \{e^{\alpha/\infty} - 1\}^{-\theta}\right)^{-\lambda} + \left(1 + \frac{1}{\delta} \{e^{\alpha/0} - 1\}^{-\theta}\right)^{-\lambda} \\
 &= -\left(1 + \frac{1}{\delta} \{e^0 - 1\}^{-\theta}\right)^{-\lambda} + \left(1 + \frac{1}{\delta} \{e^{\infty} - 1\}^{-\theta}\right)^{-\lambda} \\
 &= -\left(1 + \frac{1}{\delta} \{1 - 1\}^{-\theta}\right)^{-\lambda} + \left(1 + \frac{1}{\delta} \{\infty - 1\}^{-\theta}\right)^{-\lambda} \\
 &= -\left(1 + \frac{1}{\delta(0)^\theta}\right)^{-\lambda} + \left(1 + \frac{1}{\delta(\infty)^\theta}\right)^{-\lambda} = \\
 &-(\infty)^{-\lambda} + (1+0)^{-\lambda} = -\frac{1}{(\infty)^\lambda} + 1 = 0 + 1 = 1 \text{ verified}
 \end{aligned}$$

(4) Derivations of Survival function /Reliability function

$$R(x) = 1 - F(x) = 1 - \left[1 - \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x} - 1)^{-\theta} \right\}^{-\lambda} \right] = \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x} - 1)^{-\theta} \right\}^{-\lambda}$$

(5) Derivation of Quantile function

The quantile function is

$$Q(u) = F^{-1}(u)$$

Let $F(x) = u$. Then, $x = F^{-1}(u)$ (1)

That is, $1 - \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x} - 1)^{-\theta} \right\}^{-\lambda} = u, 0 < u < 1$.

$$\text{or, } \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x} - 1)^{-\theta} \right\}^{-\lambda} = 1 - u \text{ or, } 1 + \frac{1}{\delta} (e^{\alpha/x} - 1)^{-\theta} = (1 - u)^{-1/\lambda}$$

$$\text{or, } \frac{1}{\delta} (e^{\alpha/x} - 1)^{-\theta} = (1 - u)^{-1/\lambda} - 1, \text{ or, } (e^{\alpha/x} - 1)^{-\theta} = \delta \left[(1 - u)^{-1/\lambda} - 1 \right]$$

$$\text{or, } e^{\alpha/x} - 1 = \left[\delta \left\{ (1 - u)^{-1/\lambda} - 1 \right\} \right]^{-1/\theta} = e^{\alpha/x} = 1 + \left[\delta \left\{ (1 - u)^{-1/\lambda} - 1 \right\} \right]^{-1/\theta}$$

$$\text{or, } \frac{\alpha}{x} = \log \left[1 + \left[\delta \left\{ (1 - u)^{-1/\lambda} - 1 \right\} \right]^{-1/\theta} \right]$$

$$\text{or, } x = \alpha \left[\log \left[1 + \left[\delta \left\{ (1 - u)^{-1/\lambda} - 1 \right\} \right]^{-1/\theta} \right] \right]^{-1} \tag{2}$$

$$\text{Hence, } Q(u) = F^{-1}(u) = \alpha \left[\log \left[1 + \left[\delta \left\{ (1 - u)^{-1/\lambda} - 1 \right\} \right]^{-1/\theta} \right] \right]^{-1} \tag{3}$$

(6) Derivation of Median

For median, taking $u = 1/2$ in equation (3), we get

$$\text{Median} = \alpha \left[\log \left[1 + \left[\delta \left\{ (1/2)^{-1/\lambda} - 1 \right\} \right]^{-1/\theta} \right] \right]^{-1}$$

(7) **Derivation of Random deviate generation**

The expression (2) is the expression for random deviate generation which helps to generate random numbers (observations).

(8) **Derivation of log likelihood function and their partial derivatives with respect to four parameters**

$$l = n \ln \alpha + n \ln \lambda + n \ln \theta - n \ln \delta - 2 \sum_{i=1}^n \ln(x_i) - (\theta + 1) \sum_{i=1}^n \ln(e^{\alpha/x_i} - 1) + \alpha \sum_{i=1}^n (1/x_i) \\ - (\lambda + 1) \sum_{i=1}^n \ln \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_i} - 1)^{-\theta} \right\}$$

Differentiating partially with respect to α , we get

$$\frac{\partial l}{\partial \alpha} = \frac{n}{\alpha} - (\theta + 1) \sum_{i=1}^n \frac{d \ln(e^{\alpha/x_i} - 1)}{d \alpha} + \sum_{i=1}^n (1/x_i) - (\lambda + 1) \sum_{i=1}^n \frac{\partial \ln \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_i} - 1)^{-\theta} \right\}}{\partial \alpha} \\ = \frac{n}{\alpha} - (\theta + 1) \sum_{i=1}^n (e^{\alpha/x_i} - 1)^{-1} \times \frac{1}{x_i} e^{\alpha/x_i} + \sum_{i=1}^n (1/x_i) \\ - (\lambda + 1) \sum_{i=1}^n \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_i} - 1)^{-\theta} \right\}^{-1} \left\{ \frac{1}{\delta} (-\theta) (e^{\alpha/x_i} - 1)^{-(\theta+1)} \times \left(\frac{1}{x_i} e^{\alpha/x_i} \right) \right\} \\ = \frac{n}{\alpha} - (\theta + 1) \sum_{i=1}^n \frac{1}{x_i} (e^{\alpha/x_i} - 1)^{-1} e^{\alpha/x_i} + \sum_{i=1}^n (1/x_i) + \frac{(\lambda + 1)\theta}{\delta} \sum_{i=1}^n \frac{e^{\alpha/x_i}}{x_i} \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_i} - 1)^{-\theta} \right\}^{-1} (e^{\alpha/x_i} - 1)^{-(\theta+1)}$$

Differentiating with respect to λ , we get

$$\frac{\partial l}{\partial \lambda} = \frac{n}{\lambda} - \sum_{i=1}^n \ln \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_i} - 1)^{-\theta} \right\}$$

Differentiating with respect to θ , we get

$$\frac{\partial l}{\partial \theta} = \frac{n}{\theta} - \sum_{i=1}^n \ln(e^{\alpha/x_i} - 1) - (\lambda + 1) \sum_{i=1}^n \frac{\partial \ln \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_i} - 1)^{-\theta} \right\}}{\partial \theta} \\ = \frac{n}{\theta} - \sum_{i=1}^n \ln(e^{\alpha/x_i} - 1) - (\lambda + 1) \sum_{i=1}^n \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_i} - 1)^{-\theta} \right\}^{-1} \frac{1}{\delta} \frac{\partial (e^{\alpha/x_i} - 1)^{-\theta}}{\partial \theta} \\ = \frac{n}{\theta} - \sum_{i=1}^n \ln(e^{\alpha/x_i} - 1) - \frac{(\lambda + 1)}{\delta} \sum_{i=1}^n \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_i} - 1)^{-\theta} \right\}^{-1} \ln(e^{\alpha/x_i} - 1) \cdot (e^{\alpha/x_i} - 1)^{-\theta} \cdot (-1) \\ = \frac{n}{\theta} - \sum_{i=1}^n \ln(e^{\alpha/x_i} - 1) + \frac{(\lambda + 1)}{\delta} \sum_{i=1}^n \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_i} - 1)^{-\theta} \right\}^{-1} (e^{\alpha/x_i} - 1)^{-\theta} \ln(e^{\alpha/x_i} - 1).$$

Differentiating with respect to δ , we get

$$\begin{aligned} \frac{\partial l}{\partial \delta} &= -\frac{n}{\delta} - (\lambda + 1) \sum_{i=1}^n \frac{\partial \ln \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_i} - 1)^{-\theta} \right\}}{\partial \delta} \\ &= -\frac{n}{\delta} - (\lambda + 1) \sum_{i=1}^n \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_i} - 1)^{-\theta} \right\}^{-1} \left\{ -\frac{1}{\delta^2} (e^{\alpha/x_i} - 1)^{-\theta} \right\} \\ &= -\frac{n}{\delta} + \frac{(\lambda + 1)}{\delta^2} \sum_{i=1}^n (e^{\alpha/x_i} - 1)^{-\theta} \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_i} - 1)^{-\theta} \right\}^{-1} \end{aligned}$$

(9) Derivation of function A and their partial derivatives with respect to four parameters

$$A(x; \alpha, \lambda, \theta, \delta) = \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right]^2$$

$$\text{Or, } A(x_{(i)}; \alpha, \lambda, \theta, \delta) = \sum_{i=1}^n \left[\left[1 - \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda} \right] - \frac{i}{n+1} \right]^2 ; x > 0, (\alpha, \lambda, \theta, \delta) > 0$$

Differentiating A with respect to α , we get

$$\begin{aligned} \frac{\partial A}{\partial \alpha} &= \frac{\partial \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right]^2}{\partial \alpha} \\ &= 2 \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right] \frac{\partial}{\partial \alpha} \left[\left[1 - \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda} \right] - \frac{i}{n+1} \right] \\ &= 2 \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right] [-(-\lambda)] \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-(\lambda+1)} \times \frac{1}{\delta} (-\theta) (e^{\alpha/x_{(i)}} - 1)^{-(\theta+1)} \times \left(\frac{1}{x_{(i)}} \right) e^{\alpha/x_{(i)}} \\ &= \frac{-2\lambda\theta}{\delta} \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right] \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-(\lambda+1)} \frac{e^{\alpha/x_{(i)}}}{x_{(i)}} (e^{\alpha/x_{(i)}} - 1)^{-(\theta+1)} \end{aligned}$$

Differentiating A with respect to λ , we get

$$\begin{aligned} \frac{\partial A}{\partial \lambda} &= \frac{\partial \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right]^2}{\partial \lambda} = \frac{\partial \sum_{i=1}^n \left[\left[1 - \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda} \right] - \frac{i}{n+1} \right]^2}{\partial \lambda} \\ &= 2 \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right] \frac{\partial}{\partial \lambda} \left[\left[1 - \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda} \right] - \frac{i}{n+1} \right] \\ &= 2 \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right] [-(-1)] \ln \left[1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right] \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda} \\ &= 2 \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right] \left[\ln \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\} \right] \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda} \end{aligned}$$

Differentiating A with respect to θ , we get

$$\begin{aligned} \frac{\partial A}{\partial \theta} &= \frac{\partial \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right]^2}{\partial \theta} = \frac{\partial \sum_{i=1}^n \left[\left[1 - \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda} \right] - \frac{i}{n+1} \right]^2}{\partial \theta} \\ &= 2 \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right] \frac{\partial}{\partial \theta} \left[\left[1 - \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda} \right] - \frac{i}{n+1} \right] \\ &= 2 \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right] [-(-\lambda)] \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-(\lambda+1)} \times \frac{\partial}{\partial \theta} \left\{ \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\} \\ &= \frac{2\lambda}{\delta} \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right] \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-(\lambda+1)} \left[(e^{\alpha/x_{(i)}} - 1)^{-\theta} \ln(e^{\alpha/x_{(i)}} - 1) \right] (-1) \\ &= -\frac{2\lambda}{\delta} \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right] \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-(\lambda+1)} \left[(e^{\alpha/x_{(i)}} - 1)^{-\theta} \ln(e^{\alpha/x_{(i)}} - 1) \right] \end{aligned}$$

Differentiating A with respect to δ , we get

$$\begin{aligned} \frac{\partial A}{\partial \delta} &= \frac{\partial \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right]^2}{\partial \delta} = \frac{\partial \sum_{i=1}^n \left[\left[1 - \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda} \right] - \frac{i}{n+1} \right]^2}{\partial \delta} \\ &= 2 \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right] \frac{\partial}{\partial \delta} \left[\left[1 - \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda} \right] - \frac{i}{n+1} \right] \\ &= 2 \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right] [(-1)(-\lambda)] \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-(\lambda+1)} \times \left(-\frac{1}{\delta^2} \right) (e^{\alpha/x_{(i)}} - 1)^{-\theta} \\ &= -\frac{2\lambda}{\delta^2} \sum_{i=1}^n \left[F(X_{(i)}) - \frac{i}{n+1} \right] \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-(\lambda+1)} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \end{aligned}$$

(1) Derivation of function Z and their partial derivatives with respect to four parameters

$$\begin{aligned} Z(X_{(i)}; \alpha, \lambda, \theta, \delta) &= \frac{1}{12n} + \sum_{i=1}^n \left[F(x_{i:n} | \alpha, \lambda, \theta, \delta) - \frac{2i-1}{2n} \right]^2 \\ &= \frac{1}{12n} + \sum_{i=1}^n \left[\left[1 - \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda} \right] - \frac{2i-1}{2n} \right]^2 \end{aligned}$$

Since $\frac{1}{12n}$ is constant having derivative zero, replacing $2/(n+1)$ in LSE by $(2n-1)/2n$, we can get the

partial derivatives $\frac{\partial Z}{\partial \alpha}$, $\frac{\partial Z}{\partial \lambda}$, $\frac{\partial Z}{\partial \theta}$ and $\frac{\partial Z}{\partial \delta}$ as

$$\begin{aligned}\frac{\partial Z}{\partial \alpha} &= \frac{-2\lambda\theta}{\delta} \sum_{i=1}^n \left[F(X_{(i)}) - \frac{2i-1}{2n} \right] \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda+1} \frac{e^{\alpha/x_{(i)}}}{x_{(i)}} (e^{\alpha/x_{(i)}} - 1)^{-(\theta+1)} \\ \frac{\partial Z}{\partial \lambda} &= 2 \sum_{i=1}^n \left[F(X_{(i)}) - \frac{2i-1}{n} \right] \left[\ln \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\} \right] \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda} \\ \frac{\partial Z}{\partial \theta} &= -\frac{2\lambda}{\delta} \sum_{i=1}^n \left[F(X_{(i)}) - \frac{2i-1}{2n} \right] \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda+1} \left[(e^{\alpha/x_{(i)}} - 1)^{-\theta} \ln(e^{\alpha/x_{(i)}} - 1) \right] \\ \frac{\partial Z}{\partial \delta} &= -\frac{2\lambda}{\delta^2} \sum_{i=1}^n \left[F(X_{(i)}) - \frac{2i-1}{2n} \right] \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x_{(i)}} - 1)^{-\theta} \right\}^{-\lambda+1} (e^{\alpha/x_{(i)}} - 1)^{-\theta}\end{aligned}$$

(2) Models considered for comparison

Generalized Rayleigh (GR) distribution:

$$f_{GR}(x) = 2\alpha\lambda^2 x e^{-(\lambda x)^2} \left\{ 1 - e^{-(\lambda x)^2} \right\}^{\alpha-1}; \quad x > 0, (\alpha, \lambda) > 0$$

Exponentiated Power Lindley (EPL) distribution:

$$f_{EPL}(x) = \frac{\alpha\beta\theta^2 x^{\beta-1}}{(\theta+1)} (1+x^\beta) e^{-\theta x^\beta} \left[1 - \left(1 + \frac{\theta x^\beta}{\theta+1} \right) e^{-\theta x^\beta} \right]^{\alpha-1}; \quad x > 0, ,$$

Generalized Weibull Extension (GWE):

$$f_{GWE}(x) = \alpha \beta (\lambda x)^{\beta-1} \exp \left\{ (\lambda x)^\beta + \frac{1}{\lambda} \left(1 - \exp \left((\lambda x)^\beta \right) \right) \right\} \left[1 - \exp \left\{ \frac{1}{\lambda} \left(1 - \exp \left((\lambda x)^\beta \right) \right) \right\} \right]^{\alpha-1}; \quad x \geq 0, (\alpha, \beta, \lambda) > 0$$

Exponentiated Weibull Distribution (EW):

$$f_{EW}(x) = \alpha\beta\lambda x^{\beta-1} \exp(-\alpha x^\beta) \left(1 - \exp(-\alpha x^\beta) \right)^{\lambda-1}; \quad x \geq 0, (\alpha, \beta, \lambda) > 0$$

Generalized Exponential (GE) distribution:

$$f_{GE}(x) = \alpha\lambda e^{-\lambda x} \left\{ 1 - e^{-\lambda x} \right\}^{\alpha-1}; \quad \alpha > 0, \lambda > 0, x \geq 0$$

Half Logistic Nadarajah Haghighi (HLNHE) Distribution:

$$f_{HLNHE}(x) = \frac{2\alpha\beta\lambda(1+\alpha x)^{(\beta-1)} \exp\left(\lambda\left(1-(1+\alpha x)^\beta\right)\right)}{\left[1 + \exp\left(\lambda\left(1-(1+\alpha x)^\beta\right)\right)\right]^2}; \quad \alpha, \beta, \lambda > 0, x > 0$$

Lindley Inverse Weibull (LIW):

$$f_{LIW}(x) = \alpha\beta \left(\frac{\theta^2}{\theta+1}\right) x^{-(\beta+1)} \exp(-\alpha x^{-\beta}) \left(1 - \exp(-\alpha x^{-\beta})\right)^{\alpha-1} \left\{1 - \log\left(1 - \exp(-\alpha x^{-\beta})\right)\right\} \quad ; (\alpha, \beta, \theta) > 0, x > 0$$

Exponentiated Generalized inverted Exponential (EGIE):

$$f_{EGIE}(x) = \frac{\alpha\beta\lambda}{x^2} e^{-\lambda/x} \left\{1 - e^{-\lambda/x}\right\}^{\alpha-1} \left[1 - \left\{1 - e^{-\lambda/x}\right\}^\alpha\right]^{\beta-1} ; \alpha, \beta, \lambda > 0, x > 0$$

Lomax Exponentiated Weibull (LEW):

$$f_{LEW}(x) = \alpha\beta\theta x^{\beta-1} \left(1 - e^{-x^\beta}\right)^{\alpha-1} \left[1 - \left(1 - e^{-x^\beta}\right)^\alpha\right]^2 \left[1 + \left(1 - e^{-x^\beta}\right)^\alpha \left\{1 - \left(1 - e^{-x^\beta}\right)^\alpha\right\}^{-1}\right]^{-(\theta+1)} ; (\alpha, \beta, \theta) > 0, x > 0$$

Generalized Inverted Generalized Exponential (GIGE):

$$f_{GIGE}(x) = \alpha\lambda\gamma x^{-2} e^{-\gamma(\lambda/x)} ; \left(1 - e^{-\gamma(\lambda/x)}\right)^{\alpha-1} ; \alpha, \lambda, \gamma > 0, x > 0,$$

(10) Detailed calculation for mode

$$f(x) = \begin{cases} \frac{\lambda\theta\alpha}{\delta} x^{-2} e^{\alpha/x} \left(e^{\alpha/x} - 1\right)^{-(\theta+1)} \left[1 + \frac{1}{\delta} \left(e^{\alpha/x} - 1\right)^{-\theta}\right]^{-(\lambda+1)} ; x \geq 0, (\alpha, \lambda, \theta, \delta) > 0 \\ 0 ; \text{Otherwise} \end{cases}$$

log density function is

Differentiating with respect to x, we get

$$\ln f(x) = \ln \alpha + \ln \lambda + \ln \theta - \ln \delta - 2 \ln x + \alpha x^{-1} - (\theta + 1) \ln(e^{\alpha/x} - 1) - (\lambda + 1) \ln \left\{1 + \frac{1}{\delta} (e^{\alpha/x} - 1)^{-\theta}\right\} \quad (15)$$

$$\frac{d \ln f(x)}{dx} = \frac{d}{dx} \left[\ln f(x) = \ln \alpha + \ln \lambda + \ln \theta - \ln \delta - 2 \ln x + \right. \\ \left. \alpha x^{-1} - (\theta + 1) \ln(e^{\alpha/x} - 1) - (\lambda + 1) \ln \left\{ 1 + \frac{1}{\delta} (e^{\alpha/x} - 1)^{-\theta} \right\} \right]$$

$$\frac{f'(x)}{f(x)} = -\frac{2}{x} - \frac{\alpha}{x^2} + \frac{\alpha(\theta+1)e^{\alpha/x}}{x^2(e^{\alpha/x}-1)} - \frac{\alpha\theta(\lambda+1)e^{\alpha/x}(e^{\alpha/x}-1)^{-(\theta+1)}}{x^2\left\{\delta+(e^{\alpha/x}-1)^{-\theta}\right\}}$$

$$f'(x) = \left[-\frac{2}{x} - \frac{\alpha}{x^2} + \frac{\alpha(\theta+1)e^{\alpha/x}}{x^2(e^{\alpha/x}-1)} - \frac{\alpha\theta(\lambda+1)e^{\alpha/x}(e^{\alpha/x}-1)^{-(\theta+1)}}{x^2\left\{\delta+(e^{\alpha/x}-1)^{-\theta}\right\}} \right] f(x)$$

applying $f'(x) = 0$

$$\left[-\frac{2}{x} - \frac{\alpha}{x^2} + \frac{\alpha(\theta+1)e^{\alpha/x}}{x^2(e^{\alpha/x}-1)} - \frac{\alpha\theta(\lambda+1)e^{\alpha/x}(e^{\alpha/x}-1)^{-(\theta+1)}}{x^2\left\{\delta+(e^{\alpha/x}-1)^{-\theta}\right\}} \right] = 0, \text{ Since } f(x) \neq 0$$

$$\frac{\alpha\theta(\lambda+1)e^{\alpha/x}(e^{\alpha/x}-1)^{-(\theta+1)}}{x\left\{\delta+(e^{\alpha/x}-1)^{-\theta}\right\}} - \frac{\alpha(\theta+1)e^{\alpha/x}}{x(e^{\alpha/x}-1)} + \frac{\alpha}{x} + 2 = 0$$

Solving above equation, mode can be calculated.

Since this is non linear so it needs Newton-Raphson method for solution.